## Quantile Regression

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## Average derivative estimation (1)

- In mean regression, we focus on estimating the conditional mean

$$
\mu(\mathbf{x})=E(Y \mid \mathbf{x})
$$

- In linear regression, the partial derivatives $\partial \mu(\dot{\mathbf{x}}) / \partial x_{i}$ are assumed to be constant.
- They are of primary interest because they measure how much the mean response change as the ith covariate perturb one unit.


## Average derivative estimation (2) Rich got richer, poor got poorer?



Share of capital income earned by top 1\% and bottom 80\%, 1979-2003 (Shapiro and Friedman 2006)

## Average derivative estimation (3)

- For quantile regression, we may consider the average gradient

$$
\beta_{\alpha}=\left(\beta_{\alpha 1}, \ldots, \beta_{\alpha d}\right)=E\left(\nabla \theta_{\alpha}(\mathbf{X})\right)
$$

where $\theta_{\alpha}(\mathbf{X})$ is the conditional $\alpha$ th of Y given $\mathbf{X}$

## Average derivative estimation (4)

- Consider the model
$Y=\mu(\mathbf{X})+\tau[\mu(\mathbf{X})]^{\lambda} \varepsilon$
$\varepsilon$ and $\mathbf{X}$ are independent
$\varepsilon$ has continuous distribution function $F_{\varepsilon}$
the mean of $\varepsilon$ is zero
$\tau$ and $\lambda$ are real parameters


## Average derivative estimation (5)

- Let $e_{\alpha}$ be the $\alpha$ th quantile of $F_{\varepsilon}$

$$
\begin{aligned}
\theta_{\alpha}(\mathbf{x}) & =\mu(\mathbf{x})+\tau[\mu(\mathbf{x})]^{\lambda} e_{\alpha} \\
\nabla \theta_{\alpha}(\mathbf{x}) & =\nabla \mu(\mathbf{x})+\tau \lambda[\mu(\mathbf{x})]^{\lambda-1} \nabla_{\mu}(\mathbf{x}) e_{\alpha} \\
\beta_{\alpha} & =E(\nabla \mu(\mathbf{X}))+\tau \lambda E\left\{[\mu(\mathbf{X})]^{\lambda-1} \nabla \mu(\mathbf{X})\right\} e_{\alpha}
\end{aligned}
$$

- If assuming $F_{\varepsilon}=\Phi$

$$
\begin{aligned}
& d=\tau=\lambda=1 \\
& \mu(x)=\gamma_{1}+\gamma_{2} x
\end{aligned}
$$

then

$$
\beta_{\alpha}=\left[1+\Phi^{-1}(\alpha)\right] \gamma_{2}
$$

$$
\beta_{0.1}=-0.282 \gamma_{2}, \quad \beta_{0.5}=\gamma_{2}, \quad \beta_{0.9}=2.282 \gamma_{2}
$$

## Average derivative estimation (6)

- Taking derivative on both side of the above equation and taking expectation, we get

$$
\begin{aligned}
E\left(\omega(\boldsymbol{X}) \nabla \theta_{\alpha}(\boldsymbol{X})\right) & =\left[\int g^{\prime}\left(\gamma^{t} \boldsymbol{x}\right) \omega(\boldsymbol{x}) f(\boldsymbol{x}) d \boldsymbol{x}\right] \gamma \\
& =\beta \gamma
\end{aligned}
$$

## Average derivative estimation (7)

- We may estimate

$$
E\left(\omega(\boldsymbol{X}) \nabla \theta_{\alpha}(\boldsymbol{X})\right)=\int \nabla \theta_{\alpha}(\boldsymbol{X}) \omega(\boldsymbol{x}) f(\boldsymbol{x}) d \boldsymbol{x}
$$

by $\hat{\beta}_{1}=n^{-1} \sum\left\{\nabla \hat{\theta}\left(\mathbf{X}_{i}\right)\right\} w\left(\mathbf{X}_{i}\right)$,
where $\nabla \hat{\theta}\left(\mathbf{X}_{i}\right)$ is a nonparametric estimator of the gradient of the conditional quantile $\theta(\mathbf{x})$

## Average derivative estimation (8)

- By integration by parts,

$$
\begin{aligned}
E\left(\omega(\boldsymbol{X}) \nabla \theta_{\alpha}(\boldsymbol{X})\right) & =\int \nabla \theta_{\alpha}(\boldsymbol{X}) \omega(\boldsymbol{x}) f(\boldsymbol{x}) d \boldsymbol{x} \\
& =-\int \theta_{\alpha}(\boldsymbol{X}) \nabla\{\omega(\boldsymbol{x}) f(\boldsymbol{x})\} d \boldsymbol{x}
\end{aligned}
$$

- An alternative estimator is

$$
\begin{aligned}
\hat{\beta}_{2} & =-\frac{1}{n} \sum_{i=1}^{n} \hat{\theta}\left(\mathbf{X}_{i}\right) \frac{\nabla w\left(\mathbf{X}_{i}\right) \hat{f}\left(\mathbf{X}_{i}\right)+w\left(\mathbf{X}_{i}\right) \nabla \hat{f}\left(\mathbf{X}_{i}\right)}{\hat{f}\left(\mathbf{X}_{i}\right)} \\
& =-\frac{1}{n} \sum_{i=1}^{n} \hat{\theta}\left(\mathbf{X}_{i}\right)\left\{\nabla w\left(\mathbf{X}_{i}\right)+w\left(\mathbf{X}_{i}\right) \hat{\ell}\left(\mathbf{X}_{i}\right)\right\},
\end{aligned}
$$

where $\hat{\ell}\left(\mathbf{X}_{i}\right)=\nabla \hat{f}\left(\mathbf{X}_{i}\right) / \hat{f}\left(\mathbf{X}_{i}\right), \hat{f}$ and $\nabla \hat{f}$ are nonparametric estimator of the density and its derivative

## Average derivative estimation (9)

- Leave one out estimator

$$
\begin{aligned}
\hat{f}\left(\mathbf{X}_{i}\right) & =\frac{1}{(n-1) h_{n}^{d}} \sum_{j \neq i} W\left(\frac{\mathbf{X}_{j}-\mathbf{X}_{i}}{h_{n}}\right), \\
\nabla \hat{f}\left(\mathbf{X}_{i}\right) & =\frac{1}{(n-1) h_{n}^{d+1}} \sum_{j \neq i} W^{(1)}\left(\frac{\mathbf{X}_{j}-\mathbf{X}_{i}}{h_{n}}\right),
\end{aligned}
$$

## Penalized univariate smoothing method (1)

- Assume that

$$
y=f(x)+\epsilon
$$

- In case of the mean regression, we may consider estimate the function fbv minimizing

$$
\operatorname{RSS}(f, \lambda)=\sum_{i=1}^{N}\left\{y_{i}-f\left(x_{i}\right)\right\}^{2}+\lambda \int\left\{f^{\prime \prime}(t)\right\}^{2} d t,
$$

$\lambda=0: f$ can be any function that interpolates the data.
$\lambda=\infty$ : the simple least squares line fit, since no second derivative can be tolerated.

- It can be shown that for $\lambda \in(0, \infty)$, there is a unique minimizer, which is a natural cubic spline with knots at $x_{i}$


## Splines (1)

- Suppose that we can to estimate a function f(.) with a piecewise polynomial function
- In the simplest case, we can just estimate $f$ by a piecewise constant function
- We may write

$$
\begin{aligned}
& f(X)=\sum_{m=1}^{3} \beta_{m} h_{m}(X) \\
& h_{1}(X)=I\left(X<\xi_{1}\right) \\
& h_{2}(X)=I\left(\xi_{1} \leq X<\xi_{2}\right) \\
& h_{3}(X)=I\left(\xi_{2} \leq X\right)
\end{aligned}
$$



## Splines (2)

- We may also fit f by a piecewise linear function, which requires 3 additional base function

$$
h_{m+3}=h_{m}(X) X, m=1, \ldots, 3
$$

Piecewise Linear


## Splines (3)

- If we require continuity on the knots, we put some constraints ( 2 constraints in the above exmple) on the coefficient of

$$
f(X)=\sum_{m=1}^{M} \beta_{m} h_{m}(X),
$$

- More conveniently, we may write the base functions



## Splines (4)



Continuous First Derivative


Continuous


Continuous Second Derivative


Piecewise
cubic polynomial fit
$h_{1}(X)=1$,
$h_{2}(X)=X$,
$h_{3}(X)=X^{2}$
$h_{4}(X)=X^{3}$
$h_{5}(X)=\left(X-\xi_{1}\right)_{+}^{3}$,
$h_{6}(X)=\left(X-\xi_{2}\right)_{+}^{3}$.

## Splines (5)

- An order-M spline with knots $\xi_{j}, j=1, \ldots, K$ is a piecewise-polynomial of order M (degree M-1), has continuous derivative up to order M-2
- A cubic spline has order $\mathrm{M}=4$
- The bases are

$$
\begin{aligned}
h_{j}(X) & =X^{j-1}, j=1, \ldots, M, \\
h_{M+\ell}(X) & =\left(X-\xi_{\ell}\right)_{+}^{M-1}, \ell=1, \ldots, K .
\end{aligned}
$$

## Natural cubic splines

- In addition to requiring the function to have continuous derivatives on the knots, we require the function is linear beyond the boundary points
- If the function is represented as

$$
f(X)=\sum_{j=0}^{3} \beta_{j} X^{j}+\sum_{k=1}^{K} \theta_{k}\left(X-\xi_{k}\right)_{+}^{3} .
$$

it can be shown that

$$
\begin{array}{lc}
\beta_{2}=0, & \sum_{k=1}^{K} \theta_{k}=0 \\
\beta_{3}=0, & \sum_{k=1}^{K} \xi_{k} \theta_{k}=0
\end{array}
$$

## B-splines (1)

Let $\xi_{0}<\xi_{1}$ and $\xi_{K}<\xi_{K+1}$ be two boundary knots.

$$
\begin{aligned}
& \tau_{1} \leq \tau_{2} \leq \cdots \leq \tau_{M} \leq \xi_{0} \\
& \tau_{j+M}=\xi_{j}, j=1, \cdots, K \\
& \xi_{K+1} \leq \tau_{K+M+1} \leq \tau_{K+M+2} \leq \cdots \leq \tau_{K+2 M}
\end{aligned}
$$

$$
B_{i, 1}(x)= \begin{cases}1 & \text { if } \tau_{i} \leq x<\tau_{i+1} \\ 0 & \text { otherwise }\end{cases}
$$

$$
\text { for } i=1, \ldots, K+2 M-1
$$

$$
B_{i, m}(x)=\frac{x-\tau_{i}}{\tau_{i+m-1}-\tau_{i}} B_{i, m-1}(x)+\frac{\tau_{i+m}-x}{\tau_{i+m}-\tau_{i+1}} B_{i+1, m-1}(x)
$$

for $i=1, \ldots, K+2 M-m$.

## B-splines (2)



FIGURE 5.20. The sequence of B-splines up to order four with ten knots evenly spaced from 0 to 1 . The B-splines have local support; they are nonzero on an interval spanned by $M+1$ knots.

## Quantile regression-penalized method (1)

- For quantile regression, we may consider

$$
\min _{g \in \mathcal{G}} \sum_{i=1}^{n} \rho_{\tau}\left(y_{i}-g\left(x_{i}\right)\right)+\lambda \int\left(g^{\prime \prime}(x)\right)^{2} d x
$$

- The solution is also a natural cubic spline


## Quantile regression-penalized method (2)

Koenker, Ng, and Portnoy (1994) consider other $L_{p}$ penalties

$$
J(g)=\left\|g^{\prime \prime}\right\|_{p}=\left(\int\left(g^{\prime \prime}(x)\right)^{p}\right)^{1 / p}
$$

For $p=1$
$\min \sum_{i=1}^{n} \rho_{\tau}\left(y_{i}-g\left(x_{i}\right)\right)+\lambda \int\left|g^{\prime \prime}(x)\right| d x$

## Quantile regression-penalized method (3)

- Another way to penalize the objective function is

$$
P(g)=V\left(g^{\prime}\right)
$$

- The total variation is defined as

$$
V(f)=\sup \sum_{i=1}^{n}\left|f\left(x_{i+1}\right)-f\left(x_{i}\right)\right|,
$$

where the sup is taken over all partitions $a \leq x_{1}<\cdots<x_{n}<b$

- For absolute continuous function

$$
V(f)=\int_{a}^{b}\left|f^{\prime}(x)\right| d x
$$

## Quantile regression-penalized method (4)

- The new penalized objective function is

$$
\min _{\mathfrak{g} \in \mathcal{G}} \sum_{i=1}^{n} \rho_{\tau}\left(y_{i}-\mathfrak{g}\left(x_{i}\right)\right)+\lambda V\left(g^{\prime}\right)
$$

- The solution is a piecewise linear function


## Animal weight VS running speed


h: hoppers
s: specials including sloth, porcupine, hippopotamus

